

Cox Proportional Hazards regression and time dependent covariates

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Overview – intro for survival analysis

- •Example of survival analysis
- •Data on survival
- •Lexis diagrams and study design
- Survival function, densities and hazard rates
- •Kaplan-Meier estimate of survival curve
- Log-rank test
- •Censoring vs competing risks

Survival data

Caerphilly study – description:

Follow-up study focusing on risk factors for cardiovascular diseases.

Inclusion period: July 1979 to October 1983. **Study population:** Men aged 43-61 at the start. **Primary outcomes:** Myocardial infarction (MI) or death.

End of study: February 1999.

Survival data - example

• Caerphilly study

Lexis diagram for Caerphilly study (zoomed)

7 random persons

Survival data

• Caerphilly study

Censoring vs competing risk wrt Kaplan-Meier

- Key assumption: Censored individuals have the same future risk as those remaining in the study
- This is called non-informative (right) censoring
- Can typically not be checked in the observed data
- What happens if we study time to CVD diagnosis?
- People may die before diagnosis is this censoring?
- **No -** people who died are no longer at risk of getting a CVD diagnosis
- Here death is a competing risk (but CVD is **not** ^a competing risk for death!)

Main problem with competing risk

- \bullet Cumulative risk is over-estimated
- Equivalently: Survival probability is under-estimated

Functions in survival analysis - relationships

- •Any one of the three uniquely determines the two other
- •The hazard is often taken as the fundamental quantity, since

$$
S(t) = \exp\left(-\int_0^t h(s) \, ds\right)
$$

$$
f(t) = h(t) \cdot \exp\left(-\int_0^t h(s) \, ds\right)
$$

•Implications:

$$
h(t) = \frac{f(t)}{S(t)}
$$

Or:

$$
h(t) = \frac{d}{dt} \bigl(-\ln(S(t)) \bigr)
$$

•Also you will often encounter *the integrated hazard* defined by

$$
H(t) = \int_0^t h(s) \, ds
$$

Cox regression – model specification

- Linear model for log-hazard rates $0(1)$ T $P1$ 1 1 T $P2$ 1 2
- Equivalent to

 $0(t)$ (CAP P_1 $^{X_1}\cdot(\exp\beta_2)$ $X_{\rm 2}$

- All individuals have the same shape of (log-)hazard
- Log-hazards are "shifted up or down"
- Hazard ratios are constant at any given time \rightarrow Proportional Hazards (PH) assumption
- PH assumption concerns **all follow-up time**

Hazard for a simple parametric model: Weibull

•Weibull distribution:

$$
h(t) = \alpha \lambda t^{\alpha - 1}
$$

$$
S(t) = \exp(-\lambda t^{\alpha})
$$

•Mean

$$
E(T) = \left(\frac{1}{\lambda}\right)^{\frac{1}{\alpha}} \Gamma\left(1 + \frac{1}{\alpha}\right)
$$

Where $\Gamma(x)$ is the incomplete gamma function

- λ is called the *scale* parameter, α the shape parameter
- \bullet Defining characteristic:

Hazard is monotone, either decreasing (α < 1), constant (α = 1), or increasing ($\alpha > 1$)

Remember: This is true all the way from zero to infinity!

Weibull distribution: survival functions

Weibull distribution: hazard functions

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Weibull distribution: survival functions

Weibull distribution: hazard functions

Weibull distribution: hazard functions

Checking PH assumption

- Time in study as time scale
- We estimate HR for smoking at study entry with respect to death

```
> coxph(Surv(os_dur, status) \sim cursmoker, data = caerphilly_dat)
ca11:\cosh(formula = Surv(os_dur, status) \sim cursmoker, data = carphilly_dat)coef exp(coef) se(coef) z
cursmokerYes 0.58696
                      1.79852 0.09398 6.245 4.23e-10
Likelihood ratio test=41.24 on 1 df, p=1.347e-10
n= 1786, number of events= 516
```
Checking PH assumption

- Time in study as time scale
- We estimate $HR \sqrt{1 + \frac{1}{\sqrt{1 + \left(1 + \frac{1}{$ to death

Years

 $No - Yes$

 10.0

Checking PH assumption – model based (I)

```
> # Note: We use a binary numeric variable for smoking, not a factor
> # Different HR before or after 5 yrs?
> coxph(Surv(os_dur, status) ~ binsmoker + tt(binsmoker) + ns(agein, df = 2),
       data = carphilly_data,
+tt = function(x, t, ...) x * (t + 5)++ )
ca11:\cosph(formula = Surv(os_dur, status) \sim binsmoker + tt(binsmoker) +ns(agein, df = 2), data = caerphilly_dat, tt = function(x,
   t. ...) x * (t + 5)\text{coef} exp(coef) se(coef) z
                                                         p
binsmoker
                 0.09283 1.09727 0.30347 0.306 0.760
tt(binsmoker) 0.02985 1.03030 0.01899 1.572 0.116
ns(agein, df = 2)1 2.08454 8.04089 0.38362 5.434 5.52e-08
ns(agein, df = 2)2 1.21326 3.36442 0.15175 7.995 1.29e-15
Likelihood ratio test=134.4 on 4 df, p=< 2.2e-16n = 1786, number of events= 516
```
Checking PH assumption – model based (II)

```
> # Effect changes log-linearly over time
> coxph(Surv(os_dur, status) ~ binsmoker + tt(binsmoker) + ns(agein, df = 2),
       data = carphilly_data,+ -tt =function(x, t, ...) x * log(t)+ -+)
ca11:cosh(formula = Surv(os_dur, status) ~ binsmoker + tt(binsmoker) +
   ns(agein, df = 2), data = caerphilly_dat, tt = function(x,
   t, ...) x * log(t)coef exp(coef) se(coef) z
                                                      р
binsmoker
                           1.1825 0.2386 0.703 0.4823
               0.1677
tt(binsmoker) 0.1815 1.1990 0.1050 1.729 0.0838
ns(agein, df = 2)1 2.0838 8.0346 0.3836 5.432 5.57e-08
ns(agein, df = 2)2 1.2135 3.3652 0.1518 7.997 1.28e-15
Likelihood ratio test=134.9 on 4 df, p=< 2.2e-16n= 1786, number of events= 516
```
Thanks for your attention – questions welcome!

(Djursland, July 2015 – H Støvring)